

# STAT 650 Homework 1

Instructors: Drs. Dennis Cox and Rudolf Riedi

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Bring to class or hand in to Dr. Riedi, Duncan Hall 2082

1. Let  $U$  be a discrete random variable with values  $\{u_i\}_{i \in I}$ . Let  $A_i := \{\omega : U(\omega) = u_i\}$ . Note that the sets  $A_i$  form a partition of  $\Omega$ , and that  $\sigma(U) = \sigma(A_i, i \in I)$ . For ease of notation we write  $\mathbb{E}[Y|U] = \mathbb{E}[Y|\sigma(U)]$ .

Let  $Y$  be in  $L^1$ . Recall that  $\mathbb{E}[Y|U]$  is uniquely determined (in the sense of a.s. equality) by two properties, i.e., a r.v.  $Z$  is a.s. equal to  $\mathbb{E}[Y|U]$  if and only if

- (i)  $Z$  is  $\sigma(U)$ -measurable
- (ii)  $\mathbb{E}[\mathbb{1}_A \cdot Z] = \mathbb{E}[\mathbb{1}_A \cdot Y]$  for all  $A \in \sigma(U)$

Verify the following facts using *only* the definition of a conditional expectations, i.e., the properties ((i) and ((ii) above. Thereby note that you need to verify ((ii) only for the generating sets  $A = A_i$ .

- (a) Show that there exist real numbers  $a_i$  such that

$$\mathbb{E}[Y|U] = \sum_{i \in I} a_i \mathbb{1}_{A_i}(\omega) \quad (1)$$

where the  $a_i$  must satisfy the condition  $\mathbb{E}[a_i \mathbb{1}_{A_i}] = \mathbb{E}[Y \mathbb{1}_{A_i}]$  and can otherwise be chosen arbitrarily (note the case  $P[A_i] = 0$ ). More precisely, set  $Z = \sum_{i \in I} a_i \mathbb{1}_{A_i}(\omega)$  and verify ((i) and ((ii).

- (b) Assume that  $Y$  is independent of  $U$ . Show that

$$\mathbb{E}[Y|U] = \mathbb{E}[Y].$$

More precisely, set  $Z = \mathbb{E}[Y]$  and verify ((i) and ((ii).

- (c) Let  $h(x, y)$  be measurable and assume that  $h(U, Y)$  lies in  $L^1$ . Due to (a) we know that the following form must hold:

$$\mathbb{E}[h(U, Y)|U] = \sum_{i \in I} \alpha_i \mathbb{1}_{A_i}(\omega).$$

Compute  $\alpha_i$ .

- (d) Let  $h$  be as above. Show the following fact. If  $Y$  and  $U$  are independent, then we have  $\alpha_i P(A_i) = \mathbb{E}[h(u_i, Y)]P(A_i)$  for all  $i$ . In other words,

$$\mathbb{E}[h(U, Y)|U] = f(U) \quad \text{where } f(x) = \mathbb{E}[h(x, Y)].$$

More precisely, set  $Z = f(U)$  and verify ((ii). You may assume that ((i) holds.

- (e) Let  $h$  be as above. Show the following fact. If  $g$  is measurable such that  $g(U)Y$  is in  $L^1$  then

$$\mathbb{E}[g(U)Y|U] = g(U)\mathbb{E}[Y|U]$$

More precisely, set  $Z = g(U)\mathbb{E}[Y|U]$  and verify ((i) and ((ii).

2. Show that  $X_t = W_t^2 - t$  is a martingale w.r.t.  $\mathcal{W}_t := \sigma(W_s : s \leq t)$ .  
Hint: Write  $W_t = W_s + (W_t - W_s)$  for  $s < t$ .

3. Notation:

Generic partition  $\pi$  of a compact interval  $[S, T]$ :  $\pi = \{t_j\}_{j=0}^N$  where  $S = t_0 \leq t_1 \leq \dots \leq t_N = T$ . The partition intervals  $I_j = [t_j, t_{j+1})$  have length  $\Delta_j = (t_{j+1} - t_j)$  and indicator functions:

$$\mathbb{1}_j(t) = \begin{cases} 1 & \text{if } t_j \leq t < t_{j+1} \\ 0 & \text{else.} \end{cases}$$

Granularity:  $\|\pi\| := \max\{\Delta_j : 0 \leq j < N\}$ .

Corresponding Brownian increments:  $\Delta W_j = W(t_{j+1}) - W(t_j)$ .

- (a) From a result in class it follows that

$$\{Z_\pi(t, \omega)\}_t := \left\{ \sum_{j=0}^{N-1} W(t_j, \omega) \mathbb{1}_j(t) \right\}_t \xrightarrow{\nu, L_2} \{W(t, \omega)\}_t \quad \text{as } \|\pi\| \rightarrow 0. \quad (2)$$

Verify this explicitly, i.e., show that  $\mathbb{E} \int_S^T |Z_\pi(t, \omega) - W(t, \omega)|^2 dt \rightarrow 0$  as  $\|\pi\| \rightarrow 0$ .

Hint: write  $\int_S^T \dots = \sum_j \int_{I_j} \dots$  and note that  $\mathbb{1}_k(t)$  vanishes identically over  $I_j$  unless  $j = k$ .

- (b) In class we saw that

$$\int_0^t W(s) dW(s) := \lim W(t_j) \Delta W_j = \frac{1}{2} (W^2(t) - t) \quad (3)$$

Generalize this result as follows. Let  $\theta \in [0, 1]$ . Then,

$$\int_0^t W(s) * dW(s) := \lim W(t_j + \theta \cdot \Delta_j) \Delta W_j = \frac{1}{2} W^2(t) + (\theta - 1/2) \cdot t \quad (4)$$

This means that the Stratonovich version of the integral (3) which is obtained by setting  $\theta = 1/2$ , becomes:

$$\int_0^t W(s) \circ dW(s) := \lim W\left(\frac{t_j + t_{j+1}}{2}\right) \Delta W_j = \frac{1}{2} W^2(t) \quad (5)$$

Also, it means that  $\mathbb{E}[\int_0^t W(s) * dW(s)] = \theta \cdot t$ , which we obtained in class for  $\theta = 0$  and  $\theta = 1$ .

Hint: Exploiting the known fact (3) from class it is enough to show that  $\int_0^t W(s) * dW(s) - \int_0^t W(s) dW(s) = \theta \cdot t$ . To this end, show that

$$\sum_j \Delta_\theta W_j \Delta W_j - \theta \cdot t \xrightarrow{L_2} 0 \quad \text{as } \|\pi\| \rightarrow 0.$$

Here we set  $\Delta_\theta W_j = W(t_j + \theta \cdot \Delta_j) - W(t_j)$ . It helps to introduce  $\Delta_\theta^* W_j = W(t_{j+1}) - W(t_j + \theta \cdot \Delta_j)$ , so that  $\Delta_\theta W_j + \Delta_\theta^* W_j = \Delta W_j$  and to exploit independence of increments over disjoint intervals. Also, write  $t = \sum_j \Delta_j$  is in class. Recall that the 4th moment of a zero-mean Gaussian variable is three times the square of the variance.