

# ELEC 533 Homework 9

Due date: In class on Wednesday, November 27, 2002

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33. The power spectral density  $S_X(\nu)$  of a wss process  $X_t$  is defined as the Fourier transform of its auto-correlation function  $R_X(\tau)$ . Show the following relations:

- (a) If  $Y_t = aX_t$  then  $R_Y(\tau) = |a|^2 R_X(\tau)$  and  $S_Y(\nu) = |a|^2 S_X(\nu)$ .
- (b) Let  $U_t$  and  $V_t$  be orthogonal processes, meaning that  $\mathbb{E}[U_t V_s] = 0$  for all  $t$  and  $s$ . Set  $W_t = U_t + V_t$ . Show that  $R_W(\tau) = R_U(\tau) + R_V(\tau)$  and conclude that  $S_W(\nu) = S_U(\nu) + S_V(\nu)$ .
- (c) Let  $X_t$  be a zero mean process, meaning that  $\mu_X(t) = 0$ . Set  $Z_t = X_t + b$ . Show that  $R_Z(\tau) = R_X(\tau) + |b|^2$  and  $S_Z(\nu) = S_X(\nu) + 2\pi|b|^2\delta(\nu)$ .
- (d) If  $A_t$  is a wss process such that  $R_A(\tau) = -R_X''(\tau)$  then  $S_A(\nu) = \nu^2 S_X(\nu)$ . (Here,  $R''$  denotes the second derivative.)

Hint: use known properties of the Fourier transform. The answers can be very short.

34. Compute the spectral density for processes with the following auto-correlation:

- (a)  $R_X(\tau) = \max(1 - |\tau|/T, 0)$ . HINT:  $R_X$  is the convolution of two extremely simple functions with extremely simple Fourier transforms.

35. Let  $\{X_t\}$  be a continuous time, but discrete valued Markov process. For simplicity we assume that the values are integers. As in the case with continuous values treated in class, we define the process via its initial distribution  $P[X_0 = k]$  at time 0 and the transition probabilities

$$P[X_t = k | X_s = n]$$

for any  $t > s$ . More precisely, the marginals are

$$P[X_t = k] = \sum_{n=1}^{\infty} P[X_t = k | X_0 = n] \cdot P[X_0 = n]$$

and the f.d.d. are

$$P[X_{t_1} = k_1, \dots, X_{t_n} = k_n] = P[X_{t_n} = k_n | X_{t_{n-1}} = k_{n-1}] \cdots P[X_{t_2} = k_2 | X_{t_1} = k_1] \cdot P[X_{t_1} = k_1]$$

As in the continuous-valued case done in class, this defines a consistent f.d.d. if and only if the Chapman-Kolmogorov equations are satisfied. Here, these Chapman-Kolmogorov equations read as:

$$\sum_{n=1}^{\infty} P[X_t = k | X_s = n] \cdot P[X_s = n | X_r = m] = P[X_t = k | X_r = m].$$

- (a) Check that the Chapman-Kolmogorov equations imply consistency. To this end, assume that the Chapman-Kolmogorov equations hold and show that for any  $j = 1, \dots, n$

$$\sum_{k_j=1}^{\infty} P[X_{t_1} = k_1, \dots, X_{t_n} = k_n] = P[X_{t_1} = k_1, \dots, X_{t_{j-1}} = k_{j-1}, X_{t_{j+1}} = k_{j+1}, \dots, X_{t_n} = k_n].$$

- (b) Consider the following particular transition probability where  $t > s$  and where  $k$  and  $n$  are positive integers:

$$P[X_t = k | X_s = n] = \begin{cases} 0 & \text{if } k < n \\ e^{-\lambda(t-s)} (\lambda(t-s))^{k-n} / (k-n)! & \text{if } k \geq n. \end{cases}$$

In other words, given that  $X_s = n$  the *increment*  $X_t - X_s$  is a Poisson random variable of mean  $\lambda(t-s)$ .

Verify that the Chapman-Kolmogorov equations are satisfied. Hint 1: recall that

$$(a+b)^n = \sum_{k=0}^n \frac{n!}{k!(n-k)!} a^k b^{n-k}.$$

Hint 2: be careful about noting when  $P[X_t = k | X_s = n]$  is zero.

- (c) With the same transition probabilities as before add the initial distribution  $P[X_0 = 0] = 1$ . Show that  $X_t$  is a Poisson variable of mean  $\lambda t$ .

Note: In this case, it is possible to show that  $X_t$  is actually a Poisson renewal process. This is not obvious, since one needs to construct arrival times  $T_k$  of events which should be spaced by independent waiting times. A first step towards this goal would be to show that the increments  $X_t - X_s$  over disjoint time intervals are independent.